# CONSISTENCY OF A UNIFORM KERNEL ESTIMATOR FOR INTENSITY OF A PERIODIC POISSON PROCESS WITH UNKNOWN PERIOD

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ABSTRACT. A uniform kernel estimator for intensity of a periodic Poisson process with unknown period is presented and a proof of its consistency is discussed. The result presented in this paper is a special case of that in [3]. The aim of discussing a uniform kernel estimator is in order to be able to present a relatively simpler proof of consistency compared to that in [3]. This is a joint work with R. Helmers and R. Zitikis.

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## 1. Introduction and main result

In this paper, a uniform kernel estimator for intensity of a periodic Poisson process with unknown period is presented and a proof of its consistency is discussed. The result presented here is a special case of that in [3] and chapter 3 of [5].

Let X be a Poisson process on  $[0, \infty)$  with (unknown) locally integrable intensity function  $\lambda$ . We assume that  $\lambda$  is a periodic function with unknown period  $\tau$ . We do not assume any parametric form of  $\lambda$ , except that it is periodic. That is, for each point  $s \in [0, \infty)$  and all  $k \in \mathbb{Z}$ , with  $\mathbb{Z}$  denotes the set of integers, we have

$$\lambda(s + k\tau) = \lambda(s). \tag{1.1}$$

Suppose that, for some  $\omega \in \Omega$ , a single realization  $X(\omega)$  of the Poisson process X defined on a probability space  $(\Omega, \mathcal{F}, \mathbf{P})$  with intensity function  $\lambda$  is observed, though only within a bounded interval [0, n]. Our goal in this paper is: (a) To study construction of a uniform kernel estimator for  $\lambda$  at a given point  $s \in [0, n]$  using only a single realization  $X(\omega)$  of the Poisson process X observed in interval [0, n]. (The

requirement  $s \in [0, n]$  can be dropped if we know the period  $\tau$ .) (b) To determine the minimal conditions for having weak convergence of this estimator.

Note that, since  $\lambda$  is a periodic function with period  $\tau$ , the problem of estimating  $\lambda$  at a given point  $s \in [0, n]$  can be reduced into a problem of estimating  $\lambda$  at a given point  $s \in [0, \tau)$ . Hence, for the rest of this paper, we assume that  $s \in [0, \tau)$ .

Note also that, the meaning of the asymptotic  $n \to \infty$  in this paper is somewhat different from the classical one. Here n does not denote our sample size, but it denotes the length of the interval of observations. The size of our samples is a random variable denoted by X([0, n]).

Let  $\hat{\tau}_n$  be any consistent estimator of the period  $\tau$ , that is,  $\hat{\tau}_n \stackrel{p}{\to} \tau$ , as  $n \to \infty$ . For example, one may use the estimators constructed in [2] or perhaps the estimator investigated by [6] and [1]. Let also  $h_n$  be a sequence of positive real numbers converging to 0, that is,

$$h_n \downarrow 0 \tag{1.2}$$

as  $n \to \infty$ . With these notations, we now define an estimator of  $\lambda(s)$  as

$$\hat{\lambda}_n(s) := \frac{\hat{\tau}_n}{n} \sum_{k=-\infty}^{\infty} \frac{1}{2h_n} X\left( [s + k\hat{\tau}_n - h_n, s + k\hat{\tau}_n - h_n] \cap [0, n] \right). \tag{1.3}$$

Let us now describe the idea behind the construction of the estimator  $\hat{\lambda}_n(s)$ . Note that, since there is only one realization of the Poisson process X available, we have to combine information about the (unknown) value of  $\lambda(s)$  from different places of the window [0, n]. For this reason, the periodicity of  $\lambda$ , that is assumption (1.1), plays a crucial role and leads to the following string of (approximate) equations

$$\lambda(s) = \frac{1}{N_n} \sum_{k=-\infty}^{\infty} \lambda(s+k\tau) \mathbf{I}\{s+k\tau \in [0,n]\}$$

$$\approx \frac{1}{N_n} \sum_{k=-\infty}^{\infty} \frac{1}{2h_n} \int_{[s+k\tau-h_n,s+k\tau+h_n]\cap[0,n]} \lambda(x) dx$$

$$= \frac{1}{N_n} \sum_{k=-\infty}^{\infty} \frac{1}{2h_n} \mathbf{E} X([s+k\tau-h_n,s+k\tau+h_n]\cap[0,n])$$

$$\approx \frac{1}{N_n} \sum_{k=-\infty}^{\infty} \frac{1}{2h_n} X([s+k\tau-h_n,s+k\tau+h_n]\cap[0,n])$$

$$\approx \frac{\tau}{n} \sum_{k=-\infty}^{\infty} \frac{1}{2h_n} X([s+k\tau-h_n,s+k\tau+h_n]\cap[0,n]), \quad (1.4)$$

where

$$N_n = \#\{k: \ s + k\tau \in [0, n]\}.$$

We note that, in order to make the first  $\approx$  in (1.4) works, we require the assumptions that s is a Lebesgue point of  $\lambda$  and (1.2) holds true. We say s is a Lebesgue point of  $\lambda$ , if we have

$$\lim_{h \downarrow 0} \frac{1}{2h} \int_{-h}^{h} |\lambda(s+x) - \lambda(s)| dx = 0$$
 (1.5)

(eg. see [7], p.107-108). Thus, from (1.4) we conclude that the quantity

$$\lambda_n(s) := \frac{\tau}{n} \sum_{k=-\infty}^{\infty} \frac{1}{2h_n} X([s + k\tau - h_n, s + k\tau + h_n] \cap [0, n]), \quad (1.6)$$

can be viewed as an estimator of  $\lambda(s)$ , provided that the period  $\tau$  is known. The estimator (1.3) is obtained by replacing  $\tau$  in (1.6) by  $\hat{\tau}_n$ .

The idea described in (1.4) and (1.6) of constructing an estimator for  $\lambda(s)$  resembles that of [4] where in a similar fashion a non-parametric estimator for an intensity function which, in addition to the periodic trend, also has a polynomial trend. In [4], just like when constructing the estimator  $\lambda_n(s)$  in (1.6), the period  $\tau$  is supposed to be known.

**Theorem 1.1.** Let the intensity function  $\lambda$  be periodic and locally integrable. Furthermore, let the bandwidth  $h_n$  be such that (1.2) holds true, and

$$nh_n \to \infty$$
 (1.7)

as  $n \to \infty$ . If

$$n|\hat{\tau}_n - \tau|/h_n \stackrel{p}{\to} 0 \tag{1.8}$$

as  $n \to \infty$ , then

$$\hat{\lambda}_n(s) \stackrel{p}{\to} \lambda(s)$$
 (1.9)

as  $n \to \infty$ , provided s is a Lebesgue point of  $\lambda$ . In other words,  $\hat{\lambda}_n(s)$  is a consistent estimator of  $\lambda(s)$ .

# 2. Proofs of Theorem 1.1

Let  $B_h(x)$  denotes the interval [x - h, x + h]. To establish Theorem 1.1, first we prove

$$\frac{1}{N_n} \sum_{k=-\infty}^{\infty} \frac{1}{2h_n} X\left(B_{h_n}(s + k\hat{\tau}_n) \cap [0, n]\right) \xrightarrow{p} \lambda(s), \tag{2.1}$$

as  $n \to \infty$ , where  $N_n = \#\{k : s + k\tau \in [0, n]\}$ . By Lemma 2.1, Lemma 2.2, and Lemma 2.3, we obtain that the quantity on the l.h.s. of (2.1) is equal to  $\lambda(s) + o_p(1)$ , as  $n \to \infty$ , which of course implies (2.1). Then, to prove (1.9), it remains to check that  $\hat{\lambda}_n(s)$  can be replaced by the quantity on the l.h.s. of (2.1), i.e. we must show that the difference between  $\hat{\lambda}_n(s)$  and the quantity on the l.h.s. of (2.1)

converges in probability to zero, as  $n \to \infty$ . To show this, first we write this difference as

$$\left(\frac{\hat{\tau}_n N_n}{n} - 1\right) \frac{1}{N_n} \sum_{k=-\infty}^{\infty} \frac{1}{2h_n} X\left(B_{h_n}(s + k\hat{\tau}_n) \cap [0, n]\right), \tag{2.2}$$

that is, the quantity on the l.h.s. of (2.1) multiplied by  $(\hat{\tau}_n N_n n^{-1} - 1)$ . Since  $\lambda(s)$  is finite, by (2.1), we have that the quantity on the l.h.s. of (2.1) is  $\mathcal{O}_p(1)$ , as  $n \to \infty$ . Hence, it remains to check that

$$\left| \frac{\hat{\tau}_n N_n}{n} - 1 \right| = o_p(1), \tag{2.3}$$

as  $n \to \infty$ . By the triangle inequality, the quantity on the l.h.s. of (2.3) does not exceed

$$\left| \frac{\hat{\tau}_n N_n}{n} - \frac{\hat{\tau}_n}{\tau} \right| + \left| \frac{\hat{\tau}_n}{\tau} - 1 \right| \le \frac{\hat{\tau}_n}{n} \left| N_n - \frac{n}{\tau} \right| + \frac{1}{\tau} \left| \hat{\tau}_n - \tau \right|. \tag{2.4}$$

Note that  $|n/\tau - N_n| \leq 1$ , and  $\hat{\tau}_n = \mathcal{O}_p(1)$ , as  $n \to \infty$  (by (1.8)). Hence, the first term on the r.h.s. of (2.4) is  $\mathcal{O}_p(n^{-1})$ , as  $n \to \infty$ . By (1.8), we have that its second term is  $o_p(n^{-1})$ , as  $n \to \infty$ . Therefore we have (2.3). This completes the proof of Theorem 1.1.

In the following lemma we shall show that we may replace the random centre  $s + k\hat{\tau}_n$  of the interval  $B_{h_n}(s + k\hat{\tau}_n)$  in (2.1) by its deterministic limit  $s + k\tau$ .

**Lemma 2.1.** Suppose  $\lambda$  is periodic (with period  $\tau$ ) and locally integrable. If, in addition, (1.2) and (1.8) are satisfied, then

$$\frac{1}{N_n} \sum_{k=-\infty}^{\infty} \frac{1}{2h_n} \left| \left\{ X \left( B_{h_n}(s + k\hat{\tau}_n) \cap [0, n] \right) - X \left( B_{h_n}(s + k\tau) \cap [0, n] \right) \right\} \right| \\
= o_p(1), \tag{2.5}$$

as  $n \to \infty$ , provided s is a Lebesgue point of  $\lambda$ .

**Proof:** First note that the difference within curly brackets on the l.h.s. of (2.5) does not exceed

$$X\left(B_{h_n}(s+k\hat{\tau}_n)\Delta B_{h_n}(s+k\tau)\cap[0,n]\right). \tag{2.6}$$

Now we notice that

$$B_{h_n-|k(\hat{\tau}_n-\tau)|}(s+k\tau) \subseteq B_{h_n}(s+k\hat{\tau}_n) \subseteq B_{h_n+|k(\hat{\tau}_n-\tau)|}(s+k\tau). \quad (2.7)$$

By (2.6) and (2.7) we have

$$|\{X (B_{h_n}(s+k\hat{\tau}_n) \cap [0,n]) - X (B_{h_n}(s+k\tau) \cap [0,n])\}|$$

$$\leq 2X (B_{h_n+|k(\hat{\tau}_n-\tau)|}(s+k\tau) \setminus B_{h_n-|k(\hat{\tau}_n-\tau)|}(s+k\tau) \cap [0,n]).$$
(2.8)

Hence, to prove (2.5), it suffices to show that

$$\frac{1}{N_n} \sum_{k=-\infty}^{\infty} \frac{1}{h_n} X \left( B_{h_n+|k(\hat{\tau}_n-\tau)|}(s+k\tau) \setminus B_{h_n-|k(\hat{\tau}_n-\tau)|}(s+k\tau) \cap [0,n] \right) 
= o_p(1),$$
(2.9)

as  $n \to \infty$ . To prove (2.9) we argue as follows. Let  $\Lambda_n$  denotes the l.h.s. of (2.9), and let also  $\epsilon > 0$  be any fixed real number. Then, for any fixed  $\delta > 0$ , we have that

$$\mathbf{P}(|\Lambda_n| \ge \epsilon) \le \mathbf{P}(\{|\Lambda_n| \ge \epsilon\} \cap \{n|\hat{\tau}_n - \tau| \le \delta h_n\}) + \mathbf{P}(n|\hat{\tau}_n - \tau| > \delta h_n).$$
(2.10)

By (1.8), the second term on the r.h.s. of (2.10) is o(1), as  $n \to \infty$ . While the first term on the r.h.s. of (2.10), does not exceed  $\mathbf{P}(|\bar{\Lambda}_n| \ge \epsilon)$ , where

$$\bar{\Lambda}_n = \frac{1}{N_n} \sum_{k=-\infty}^{\infty} \frac{1}{h_n} X\left(B_{h_n + \delta h_n}(s + k\tau) \setminus B_{h_n - \delta h_n}(s + k\tau) \cap [0, n]\right). \tag{2.11}$$

Next, by Markov inequality for the first moment, we have that

$$\mathbf{P}(|\bar{\Lambda}_n| \ge \epsilon) \le \epsilon^{-1} \mathbf{E} |\bar{\Lambda}_n|,$$

and  $\epsilon^{-1}\mathbf{E}|\bar{\Lambda}_n|$  can also be written as

$$\frac{1}{\epsilon N_n} \sum_{k=-\infty}^{\infty} \frac{1}{h_n} \int_{B_{(1+\delta)h_n}(0) \setminus B_{(1-\delta)h_n}(0)} \lambda(s+k\tau+x) \mathbf{I}(s+k\tau+x \in [0,n]) dx$$

$$= \frac{1}{\epsilon N_n} \frac{1}{h_n} \int_{B_{(1+\delta)h_n}(0) \setminus B_{(1-\delta)h_n}(0)} \lambda(s+x) \sum_{k=-\infty}^{\infty} \mathbf{I}(s+k\tau+x \in [0,n]) dx.$$
(2.12)

Now we can easily see that

$$\sum_{k=-\infty}^{\infty} \mathbf{I}(s + k\tau + x \in [0, n]) \le N_n + 1.$$

Then, the r.h.s. of (2.12) does not exceed

$$\frac{1}{\epsilon h_n} \left( \frac{1}{N_n} + 1 \right) \int_{B_{(1+\delta)h_n}(0) \setminus B_{(1-\delta)h_n}(0)} \lambda(s+x) dx. \tag{2.13}$$

We also can see that  $N_n^{-1} + 1 \le 2$ . Furthermore, the quantity in (2.13) can be bounded above by

$$\frac{2}{\epsilon h_n} \int_{B_{(1+\delta)h_n}(0)} |\lambda(s+x) - \lambda(s)| dx 
+ \frac{2}{\epsilon h_n} |B_{(1+\delta)h_n}(0) \setminus B_{(1-\delta)h_n}(0)| \lambda(s).$$
(2.14)

Since s is a Lebesgue point of  $\lambda$ , the first term of (2.14) converges to zero as  $n \to \infty$ . While the second term of (2.14) does not exceed

 $8\epsilon^{-1}\delta\lambda(s)$ . By taking  $\delta = \delta_n \downarrow 0$  as  $n \to \infty$ , we also have that this term converges to zero as  $n \to \infty$ . Then we get that  $\mathbf{P}(|\Lambda_n| \ge \epsilon) \to 0$  as  $n \to \infty$ , which is equivalent to (2.9). This completes the proof of Lemma 2.1.

To complete our proof of Theorem 1.1 we also need the following lemma.

**Lemma 2.2.** Suppose  $\lambda$  is periodic (with period  $\tau$ ) and locally integrable. If, in addition, (1.2) and (1.7) are satisfied, then

$$\frac{1}{N_n} \sum_{k=-\infty}^{\infty} \frac{1}{2h_n} |X(B_{h_n}(s+k\tau) \cap [0,n]) - \mathbf{E}X(B_{h_n}(s+k\tau) \cap [0,n])|$$

$$= o_p(1), \tag{2.15}$$

as  $n \to \infty$ , provided s is a Lebesgue point of  $\lambda$ .

**Proof:** First note that, for large n, the random variables

$$X(B_{h_n}(s+k\tau)\cap [0,n]),$$

for all  $k \in \mathbb{Z}$ , are independent. Then, by Chebyshev's inequality, to prove (2.15) its suffices to check that

$$\left(\frac{1}{2N_n h_n}\right)^2 \sum_{k=-\infty}^{\infty} Var\left\{X\left(B_{h_n}(s+k\tau) \cap [0,n]\right)\right\} = o(1), \quad (2.16)$$

as  $n \to \infty$ . Since X is a Poisson random variable,  $Var(X) = \mathbf{E}X$ , and for each k, we can write

$$\mathbf{E}X (B_{h_n}(s+k\tau) \cap [0,n]) = \int_{B_{h_n}(0)} \lambda(s+k\tau+x) \mathbf{I}(s+k\tau+x \in [0,n]) dx.$$
 (2.17)

Because  $\lambda$  is periodic (with period  $\tau$ ), we have that  $\lambda(s+k\tau+x) = \lambda(s+x)$ , and we also have that  $\sum_{k=-\infty}^{\infty} \mathbf{I}(s+k\tau+x) \in [0,n] \leq N_n+1$ . Then, to prove (2.16), its suffices to show

$$\frac{1}{2} \left( \frac{N_n + 1}{N_n} \right) \left( \frac{1}{N_n h_n} \right) \left( \frac{1}{2h_n} \int_{B_{h_n}(0)} \lambda(s + x) dx \right) = o(1), \quad (2.18)$$

as  $n \to \infty$ . Because s is a Lebesgue point of  $\lambda$ , we have

$$(2h_n)^{-1} \int_{B_{h_n}(0)} \lambda(s+x) dx = \lambda(s) + o(1),$$

as  $n \to \infty$ , which is finite. Because  $N_n h_n \to \infty$  as  $n \to \infty$ , (by (1.7)), then we get (2.18). This completes the proof of Lemma 2.2.

It remains to evaluate a non-random sum.

**Lemma 2.3.** Suppose  $\lambda$  is periodic (with period  $\tau$ ) and locally integrable. If, in addition, (1.2) is satisfied, then

$$\frac{1}{N_n} \sum_{k=-\infty}^{\infty} \frac{1}{2h_n} \mathbf{E} X \left( B_{h_n}(s+k\tau) \cap [0,n] \right) = \lambda(s) + o(1), \qquad (2.19)$$

as  $n \to \infty$ , provided s is a Lebesgue point of  $\lambda$ .

**Proof:** Using the fact that X is Poisson, the l.h.s. of (2.19) can be written as

$$\frac{1}{N_n} \sum_{k=-\infty}^{\infty} \frac{1}{2h_n} \int_{-h_n}^{h_n} \lambda(s+k\tau+x) \mathbf{I}(s+k\tau+x \in [0,n]) dx$$

$$= \frac{1}{2N_n h_n} \int_{-h_n}^{h_n} \lambda(s+x) \sum_{k=-\infty}^{\infty} \mathbf{I}(s+k\tau+x \in [0,n]) dx. \quad (2.20)$$

Now note that

$$(N_n - 1) \le \sum_{k = -\infty}^{\infty} \mathbf{I}(s + k\tau + x \in [0, n]) \le (N_n + 1),$$

which implies  $N_n^{-1} \sum_{k=-\infty}^{\infty} \mathbf{I}(s+k\tau+x \in [0,n])$  can be written as  $(1+\mathcal{O}(n^{-1}))$ , as  $n\to\infty$ , uniformly in x. Then, the quantity on the r.h.s. of (2.20) can be written as

$$\left(1 + \mathcal{O}\left(\frac{1}{n}\right)\right) \frac{1}{2h_n} \int_{-h_n}^{h_n} \lambda(s+x) dx.$$
(2.21)

By (1.2) together with the assumption that s is a Lebesgue point of  $\lambda$ , we have that  $(2h_n)^{-1} \int_{-h_n}^{h_n} \lambda(s+x) dx = \lambda(s) + o(1)$ , as  $n \to \infty$ . Then we get this lemma. This completes the proof of Lemma 2.3.

### References

- [1] M. Bebbington and R. Zitikis (2004). A robust heuristic estimator for the period of a Poisson intensity function. *Methodology and Computing in Applied Probability*, **6**, 441-462.
- [2] R. Helmers and I W. Mangku (2003), On estimating the period of a cyclic Poisson process. *Mathematical Statistics and Applications: Festschrift in honor of Constance van Eeden.* (Editors: Marc Moore, Sorana Froda and Christian Leger), IMS Lecture Notes Series Monograph Series, Volume 42.
- [3] R. Helmers, I W. Mangku, and R. Zitikis (2003), Consistent estimation of the intensity function of a cyclic Poisson process. J. Multivariate Anal. 84, 19-39.
- [4] R. Helmers and R. Zitikis (1999). On estimation of Poisson intensity functions, Ann. Inst. Stat. Math, 51, 2, 265-280.
- [5] Mangku, I W. (2001). Estimating the Intensity of a Cyclic Poisson Process, Ph.D Thesis, University of Amsterdam, The Netherlands.
- [6] D. Vere-Jones (1982). On the estimation of frequency in point-process data. J. Appl. Probab. 19A, 383-394.
- [7] R. L. Wheeden and A. Zygmund (1977), Measure and Integral: An Introduction to Real Analysis. Marcel Dekker, Inc., New York.